# A Collision-Based Model for Multi User Interference in Impulse Radio UWB Networks

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*Abstract* – Modeling Multi User Interference (MUI) in Impulse Radio (IR) – Ultra Wide Band (UWB) networks is addressed in this paper. The reference scenario consists of multiple asynchronous users transmitting IR-UWB signals using Pulse Position Modulation (PPM) in combination with Time Hopping (TH) coding. We provide a novel analytical expression for the average BER based on the observation that interference in IR is provoked by collisions occurring between pulses belonging to different transmissions. The proposed method requires specification of a similar set of system parameters as Gaussian-based approaches, but shows improved accuracy in estimating BER.

## I. INTRODUCTION

Different methods have been proposed in the recent past for evaluating the effect of Multi User Interference (MUI) on the performance of Impulse Radio-Ultra Wide Band (IR-UWB) networks. Earlier contributions ([1]-[2]) were inspired by the legacy of CDMA reference literature [3] adopting the Standard Gaussian Approximation (SGA) model for MUI. According to this model, the cumulative effect of all disturbing contributions at the receiver is assumed to be an additive Gaussian noise with uniform power spectral density over the range of frequencies of interest. Further investigations showed, however, as in the CDMA case [4], that the SGA provides weak estimations of BER when low values of user bit rate [5] or sparse topologies [6] are considered. Recent papers ([7]-[9]) propose analytical non-Gaussian approaches leading to exact BER expressions at the price of increased computational complexity. These methods do not express, however, BER in an explicit form, and therefore finding a direct relation between system parameters and performance remains an open question. Different perspectives for modeling interference in UWB networks were introduced in [10] and [11]. In [10], the UWB signal is modeled as a filtered Poisson random signal characterized by an average count rate  $\lambda$  (pulse inter-arrival time). It is shown that when  $\lambda$  is large, the interference provoked by an UWB signal tends to a Gaussian random process. In [11], the authors derive an approximation of the probability density function (pdf) for the interference noise at the output of a

2PPM-IR-UWB receiver. Such a pdf, however, is derived for a specific waveform shape, that is the second derivative of the Gaussian pulse, and under the assumptions of power control at the reference receiver, and in the absence of thermal noise. A novel perspective that explicitly takes into account the peculiar way in which information is structured in IR transmissions was introduced in [12]. In [12], MUI was modeled based on the observation that interference in IR is provoked by collisions occurring between pulses belonging to different transmissions. The probability of error over a single bit was in [12] explicitly related to the probability of experiencing pulse overlaps by more than a pre-defined threshold, that is, a pulse error was defined without introducing a receiver structure. The present work extends [12] by redefining the event of pulse error based on a complete receiver structure definition. In particular, the model for the receiver includes soft detection which produces an estimate of a current bit value by collecting information conveyed by the set of pulses representing it. In this sense, the proposed model moves away from the original scheme in which a typical hard detection was implicitly assumed. Note that pulse collision might also turn out to be constructive, that is, collision may result in an increased probability of correct bit detection. This circumstance was taken into account in [12] under the rather simplistic hypothesis that the probability of having an error on the bit given a collision is 0.5. In the present work, we introduce a refined definition for the probability of bit error given that one or more collisions have occurred, which also incorporates the effect of thermal noise in the receiver.

The paper is organized as follows. Section 2 defines the system model. Section 3 introduces the new MUI model. Section 4 presents results obtained by simulation of a network composed of a few nodes. Section 5 contains the conclusions.

## II. SYSTEM MODEL

The system model consists of a reference transmitter TX emitting IR-UWB-TH-PPM signals to a reference receiver RX. The binary sequence  $\mathbf{b}$  generated by TX is formed by

independent and identically distributed random variables with equally probable bits. The transmitted signal is:

$$s_{TX}(t) = \sqrt{E_{TX}} \sum_{j} p_0 \left( t - jT_s - \theta_j - \varepsilon b_{\lfloor j/N_s \rfloor} \right)$$
(1)

where  $p_{\theta}(t)$  is the energy-normalized waveform of the transmitted pulses,  $E_{TX}$  is the energy of each pulse,  $T_S$  is the average pulse repetition period,  $0 \le \theta < T_S$  is the time shift of the *j*-th pulse provoked by the TH code,  $\varepsilon$  is the PPM shift,  $b_x$  is the *x*-th bit of **b**,  $N_S$  is the number of pulses transmitted for each bit, and  $\lfloor x \rfloor$  is the inferior integer part of *x*. According to (1), the PPM modulator introduces a delay  $\varepsilon$  on all  $N_S$  pulses corresponding to a "1" bit.

A general flat channel model is assumed. The impulse response for the channel is given by  $h(t)=\alpha \delta(t-\tau)$ , where  $\alpha$ and  $\tau$  are the amplitude gain and propagation delay. TX and RX are assumed to be perfectly synchronized, that is, RX has perfect knowledge of  $\tau$ . The channel output is corrupted by thermal noise and MUI generated by  $N_i$  interfering IR-UWB devices. The received signal thus writes:

$$s_{RX}(t) = r_u(t) + r_{mui}(t) + n(t)$$
 (2)

where  $r_u(t)$ ,  $r_{mui}(t)$ , and n(t) are the useful signal, MUI, and thermal noise, respectively. As regards  $r_u(t)$ , one has:

$$r_{u}(t) = \sqrt{E_{u}} \sum_{j} p_{0} \left( t - jT_{s} - \theta_{j} - \varepsilon b_{\lfloor j/N_{s} \rfloor} - \tau \right)$$
(3)

where  $E_u = \alpha^2 E_{TX}$ . As regards  $r_{mul}(t)$ , we assume that all interfering signals are characterized by same  $T_s$ , and thus:

$$r_{mui}(t) = \sum_{n=1}^{N_i} \sqrt{E^{(n)}} \sum_{j} p_0 \left( t - jT_s - \theta_j^{(n)} - \varepsilon b_{\lfloor j/N_s^{(n)} \rfloor} - \tau^{(n)} \right)$$
(4)

where  $E^{(n)}$  and  $\tau^{(n)}$  are received energy per pulse and delay for the *n*-th interfering user. The relative delay  $\Delta \tau^{(n)} = \tau - \tau^{(n)}$  is assumed to be a random variable uniformly distributed between 0 and  $T_S$ . The terms  $\theta_j^{(n)}$  and  $N_S^{(n)}$  in (4) are the time shift of the *j*-th pulse and the number of pulses per bit for the *n*-th user, respectively. TH codes are randomly generated and correspond to pseudo noise sequences. n(t) in (2) is Gaussian noise, with double-sided power density  $\mathcal{N}_0/2$ .

A coherent correlator followed by a ML detector forms RX. Soft decision detection is performed, that is, the signal formed by  $N_S$  pulses is considered as a single multi-pulse signal. The received signal is thus cross-correlated with a correlation mask that is matched with the train of pulses representing one bit. The output of the correlator Z(x), for a generic bit  $b_x$ , can be thus expressed as follows:

$$Z(x) = \int_{xN_{S}T_{S}+\tau}^{(x+1)N_{S}T_{S}+\tau} s_{RX}(t) m_{x}(t-\tau) dt = Z_{u} + Z_{mui} + Z_{n}$$
(5)

where  $m_x(t)$  is the correlation mask for  $b_x$ , i.e.:

$$m_{x}(t) = \sum_{j=xN_{s}}^{(x+1)N_{s}} \left( p_{\theta} \left( t - jT_{s} - \theta_{j} \right) - p_{\theta} \left( t - jT_{s} - \theta_{j} - \varepsilon \right) \right)$$
(6)

Equation (5) indicates that the decision variable Z(x) consists of 3 terms: the signal term  $Z_u$ , the MUI contribution

 $Z_{mui}$ , and the noise contribution  $Z_n$  which is Gaussian with zero mean and variance  $\sigma_n^2 = N_S \mathcal{N}_0(1-R_0(\varepsilon))$ . Bit  $b_x$  is estimated by comparing Z(x) with a zero-valued threshold according to the following rule: when Z(x)>0 (Z(x)<0) decision is "0"("1"). For independent and equiprobable transmitted bits, the average BER is thus:

$$BER = \frac{1}{2} Prob(Z(x) < 0 | b_x = 0) + \frac{1}{2} Prob(Z(x) > 0 | b_x = 1) =$$
  
=  $Prob(Z(x) < 0 | b_x = 0)$  (7)

## III. ESTIMATING THE BER

Under the SGA hypothesis,  $Z_{mui}$  and  $Z_n$  would be both modeled as Gaussian variables with zero mean and variance  $\sigma_{mui}^2$  and  $\sigma_n^2$ , respectively. The average BER writes [13]:

$$BER = \frac{1}{2} \operatorname{erfc} \left\{ \sqrt{\frac{1}{2}} \left( \left( N_s \frac{E_u}{\mathcal{N}_0} (I - R_0(\varepsilon)) \right)^{-1} + \left( \frac{N_s T_s (I - R_0(\varepsilon))^2}{\sigma_M^2 \sum_{n=1}^{N_s} \frac{E^{(i)}}{E_u}} \right)^{-1} \right)^{-1} \right\}$$
(8)

where  $R_0(t)$  is the autocorrelation function of the pulse waveform  $p_0(t)$ , and  $\sigma_M^2$  is given by:

$$\sigma_M^2 = \int_0^{T_s} \left( \int_0^{T_s} p_0(t - \xi) (p_0(t) - p_0(t - \varepsilon)) dt \right)^2 d\xi$$
(9)

Note that the SGA derives from the central limit theorem and is thus only valid asymptotically.

We now introduce our proposed model which moves away from the SGA approach. We start by observing that the  $Z_u$  term in (5) is given by:

$$Z_{u} = \begin{cases} +N_{s}\sqrt{E_{u}}(l-R_{o}(\varepsilon)) & \text{if } b_{x}=0\\ -N_{s}\sqrt{E_{u}}(l-R_{o}(\varepsilon)) & \text{if } b_{x}=1 \end{cases}$$
(10)

according to which (7) can be rewritten as follows:

$$\mathbf{BER} = \operatorname{Prob}\left(Z_{mui} < -\left(N_S \sqrt{E_u} \left(I - R_0(\varepsilon)\right) + Z_n\right)\right) = \operatorname{Prob}\left(Z_{mui} < -y\right) \quad (11)$$

where  $y=N_S \sqrt{E_u} (l-R_0(\varepsilon))+Z_n$  is Gaussian with mean  $N_S \sqrt{E_u} (l-R_0(\varepsilon))$  and variance  $\sigma_n^2$ . We have thus:

$$\mathbf{BER} = \int_{-\infty}^{+\infty} \operatorname{Prob}(Z_{mui} < -y \mid y) p_{Y}(y) \, dy \tag{12}$$

where  $p_{Y}(y)$  is the Gaussian probability density function of y.

In our approach,  $Prob(Z_{mui} < y|y)$  takes into account collisions between pulses of different transmissions. The number of interfering pulses  $N_C$  is confined between 0 and  $N_SN_i$  given  $N_S$  pulses per bit, and  $N_i$  interfering users. For a given  $N_i$ , we can thus write:

$$Prob(Z_{mui} < -y | y) = P_{Z_{mui}}(y, 0)P_{CP}(0) + P_{Z_{mui}}(y, 1)P_{CP}(1) + \dots + P_{Z_{mui}}(y, N_i N_S)P_{CP}(N_i N_S)$$
(13)

where we have introduced  $P_{Zmui}(y,i)=Prob(Z_{mui} < -y|y,N_C=i)$ and  $P_{CP}(i)=Prob(N_C=i)$ . One obtains:

$$BER = \sum_{i=0}^{N_{g}N_{i}} P_{CP}(i) \int_{-\infty}^{+\infty} P_{Z_{mui}}(y,i) p_{Y}(y) dy$$
(14)

For independent interferers,  $P_{CP}(i)$  is:

$$P_{CP}(i) = {\binom{N_i N_s}{i}} P_{C0}{}^i (I - P_{C0})^{N_i N_s - i}$$
(15)

where  $P_{C0}$  is the probability that a single interfering device produces a colliding pulse within  $T_s$ .  $P_{C0}$  can be computed as a fraction of  $T_s$  during which the receiver may be affected by the presence of an interfering pulse and produce non-zero contributions to  $Z_{mui}$ , and can thus be expressed as follows:

$$P_{C0} = \frac{\min(2T_M + \varepsilon, 4T_M, T_S)}{T_S}$$
(16)

where  $T_M$  is the length of  $p_0(t)$ , defined as the period of time in which a given percentage of the pulse energy is contained. Equation (16) indicates that the time of possible collision is equal to the correlator window  $(2T_M + \varepsilon)$ , except when  $(2T_M + \varepsilon)$  is either 4 times greater than  $T_M$  or greater than  $T_S$ .

The next step for estimating BER is to define the shape of the probability function  $P_{Zmui}(y,i)$ . As shown in [11], the cumulative density function of MUI caused by one single interferer can be reasonably fitted by a linear function. Based on [11], we propose for  $P_{Zmui}(y,i)$  a linear model including multiple interferers with different received powers (see Fig.1).  $P_{Zmui}(y,i)$  is expressed by:

$$P_{Z_{max}}(y,i) = \begin{cases} 1 & \text{if } y \le -Z_{max}(i) \\ 1 - \frac{P_{CP}(i)}{2} \left( 1 + \frac{y}{Z_{max}(i)} \right) & \text{if } -Z_{max}(i) < y \le 0 \\ \frac{P_{CP}(i)}{2} \left( 1 - \frac{y}{Z_{max}(i)} \right) & \text{if } 0 < y \le Z_{max}(i) \\ 0 & \text{if } Z_{max}(i) < y \end{cases}$$
(17)

where  $Z_{max}(i)$  is the maximum value for  $Z_{mui}$  when *i* colliding pulses are present. For *i* collisions, the linear model in Fig. 1 indicates that an error occurs with probability 1 if the sum of  $Z_u$  and  $Z_n$ , i.e. *y*, is negative and lower than  $-Z_{max}(i)$ . Such a probability of error decreases with *y* with a linear slope depending in turn on  $P_{CP}(i)$ . Obviously, the probability of bit error is 0 when *y* is positive and higher than  $Z_{max}(i)$ . In this case, in fact, MUI does not provoke an error since  $Z_{mui}$  is below the *y* term. Power control is not assumed at RX, and therefore  $Z_{max}(i)$  varies. We suppose that:

$$Z_{max}(i) = \sum_{j=1}^{N_i} \left( \left\lceil \frac{i-j+1}{N_i} \right\rceil \sqrt{E_s^{(j)}} \right)$$
(18)

where  $E_S^{(1)}, E_S^{(2)}, ..., E_S^{(Ni)}$  are the interfering energies  $E^{(1)}, E^{(2)}, ..., E^{(Ni)}$  of (4), sorted so that  $E_S^{(j)} \ge E_S^{(j+1)}$  for  $j \in [1, N_i - 1]$ . Eq. (18) is an upper BER bound since  $Z_{max}(i)$  is estimated by privileging users with highest interfering energies. When substituting (17) into (14), and after a few simplifications, one can find the following BER upper bound:



Figure 1. Linear model for function  $P_{Zmui}(y,i)$ .

$$BER_{UPBOUND} = \frac{1}{2} \operatorname{erfc} \left( \sqrt{\frac{1}{2} \frac{N_s E_u}{\mathcal{N}_0} (1 - R_0(\varepsilon))} \right) + \sum_{i=0}^{N_i N_s} \frac{P_{CP}(i)}{2} \Omega \left( \frac{N_s E_u}{\mathcal{N}_0} (1 - R_0(\varepsilon)), \frac{Z_{\max}(i)^2}{N_s \mathcal{N}_0 (1 - R_0(\varepsilon))} \right)$$
(19)

where

$$\Omega(A,B) = \frac{1}{2} \operatorname{erfc}\left(\sqrt{\frac{A}{2}} - \sqrt{\frac{B}{2}}\right) + \frac{1}{2} \operatorname{erfc}\left(\sqrt{\frac{A}{2}} + \sqrt{\frac{B}{2}}\right) - \operatorname{erfc}\left(\sqrt{\frac{A}{2}}\right)$$

BER<sub>UPBOUND</sub> as in (19) includes a first term of thermal noise that only depends on signal to thermal noise ratio at RX input, and a second term accounting for MUI. Note that for computing (19), no additional information with respect to the BER computation under the SGA is requested.

### IV. SIMULATION RESULTS

Simulation of a network of 4 nodes provided the results presented in Figs.2 and 3 for two reference signal formats. In both cases, 4 users are considered  $(N_i=3)$ , and power control is assumed at RX. In the case of Fig. 2, transmitted signals have  $N_S = 2$  and  $T_S = 25$ ns, leading to  $R_b = 20$  Mb/s (signal format A). In Fig. 3, transmitted signals  $N_S = 4$  and  $T_s = 25$ ns, leading to  $R_b = 10$  Mb/s (signal format *B*). In both cases,  $p_0(t)$  is the second derivative Gaussian waveform [13], with  $T_M = 1$  ns and  $\varepsilon = 1$  ns. Performance is expressed by BER vs. signal to noise ratio  $E_b/\mathcal{N}_0$ , where  $E_b=N_SE_u$  is the received energy per bit. Note that Figs. 2 and 3 are computed for same  $E_b$ , meaning that  $E_u$  is different in the two figures. Thus, one should not be surprised if performance seems to degrade from  $N_S = 2$  to  $N_S = 4$  since  $E_u$  is smaller for  $N_S = 4$ . BER estimates based on Pulse Collision (squares) are plotted against simulation values (solid line) and SGA values (circles). Note that Pulse Collision values very well fit simulation data, while SGA underestimates BER.

Figure 4 compares Pulse Collision vs. SGA for to signal format A and B, when increasing  $N_i$ . Observe that SGA tends to Pulse Collision for high  $N_i$ , that is, when the number of collisions at the receiver input justifies the application of the central limit theorem. Preliminary investigations obtained by removing the hypothesis of power control or by varying  $N_i$  and  $N_s$  seam to lead to similar network behavior.



Figure 2. BER vs.  $E_b/\mathcal{N}_0$  with signal format A ( $N_s = 2$ ,  $T_s = 25$  ns) and  $N_i = 3$ .



Figure 3. BER vs.  $E_b/\mathcal{N}_0$  with signal format B ( $N_s = 4$ ,  $T_s = 25$  ns) and  $N_s = 3$ .



Figure 4. Theoretical BER vs. number interfering users  $N_i$  for the two models (Pulse Collision and SGA) and for the two different signal formats (signal format *A* and *B*).

## V. CONCLUSIONS

The present work extends [12] by redefining the event of pulse error based on a complete receiver structure definition. The proposed receiver model includes soft detection. A refined definition of the bit error probability when one or more collisions have occurred, which also incorporates the effect of thermal noise in the receiver, is introduced. Results show that the proposed model provides an excellent fit of simulated performance measurements, and proves to be particularly effective in those cases where SGA fails to predict network behavior. This model does not require to specify any additional system parameters compared to SGA.

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#### REFERENCES

- R. A. Scholtz, "Multiple access with time-hopping impulse modulation," IEEE Military Communications Conference, Vol. 2, 1993, pp. 447–450.
- [2] M. Z. Win and R.A. Scholtz, "Ultra-wide bandwidth time-hopping spreadspectrum impulse radio for wireless multiple access communications," IEEE Transactions on Communications, Vol. 48, Issue 4, 2000, pp. 679-689.
- [3] M. B. Pursley, "Performance Evaluation for Phase-Coded Spread Septrum Multiple-Access Communications – Part I: System Analysis," IEEE Transactions on Communications, Vol. 25, Issue 8, 1977, pp. 795-799.
- [4] M. O. Sunay and P. J. McLane, "Calculating error probabilities for DS-CDMA systems: when not to use the Gaussian approximation," IEEE Global Telecommunications Conference 1996 (GLOBECOM '96), Vol. 3, 1996, pp. 1744 – 1749.
- [5] G. Durisi and G. Romano, "On the validity of Gaussian approximation to characterize the multiuser capacity of UWB TH PPM," IEEE Conference on Ultra Wideband Systems and Technologies, 2002, pp. 157–161.
- [6] G. Giancola, L. De Nardis, and M.-G. Di Benedetto, "Multi User Interference in Power-Unbalanced Ultra Wide Band systems: Analysis and Verification," IEEE Conference on Ultra Wideband Systems and Technologies, 2003, pp. 325-329.
- [7] B. Hu and N.C. Beaulieu, "Exact bit error rate analysis of TH-PPM UWB systems in the presence of multiple access interference," IEEE Communications Letters, Vol. 7, Issue 12, 2003, pp. 572–574.
- [8] M. Sabattini, E. Masry and L.B. Milstein, "A non-Gaussian approach to the performance analysis of UWB THBPPM systems," IEEE Conference on Ultra Wideband Systems and Technologies, 2003, pp. 52-55.
- [9] G. Durisi and S. Benedetto, "Performance evaluation of TH-PPM UWB systems in the presence of multiuser interference," IEEE Communications Letters, Vol. 7, Issue: 5, 2003, pp. 224 - 226.
- [10] J. R. Fontana, "An insight into UWB interference from a shot noise perspective," IEEE Conference on Ultra Wideband Systems and Technologies, 2002, pp. 309–313.
- [11] A. R. Forouzan, M. Nasiri-Kenari, and J. A. Salehi, "Performance analysis of time-hopping spread-spectrum multiple-access systems: uncoded and coded schemes," IEEE Transactions on Communications, Vol. 1, Issue: 4, 2002, pp. 671-681.
- [12] M.-G. Di Benedetto, L. De Nardis, M. Junk, and G. Giancola, "(UWB)^2: Uncoordinated, Wireless, Baseborn, medium access control for UWB communication networks," to appear in Mobile Networks and Applications special issue on WLAN Optimization at the MAC and Network Levels (2005).
- [13] M.-G. Di Benedetto and G. Giancola, "Understanding ultra wide band radio fundamentals," Prentice Hall, 2004.